

To Pool or Not To Pool: Analyzing the Regularizing Effects of Group-Fair Training on Shared Models



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Fairness and Overfitting

Given per-group samples $(\boldsymbol{x}, \boldsymbol{y}) = \boldsymbol{z}_i \in (\mathcal{X} \times \mathcal{Y})^{\boldsymbol{m}_i}$ from $\mathcal{D}_i^{\boldsymbol{m}_i}$ assume hypothesis class $\mathcal{H} \subseteq \mathcal{X} \to \mathcal{Y}'$, loss function $\ell: \mathcal{Y}' \times \mathcal{Y} \to \mathbb{R}$

- Subgroups have potentially different distributions $\mathcal{D}_{1:q}$
- Worst-case generalization error can be analyzed on a per-group basis, but:
 - There may **too few samples** to compute bounds for minority groups
 - Data from large groups may **regularize overfitting** to small groups
- Our goal: Use majority data to bound minority overfitting

Setting: Malfare-based Fair Machine Learning —

Empirical per-group risk:

$$\hat{\mathrm{R}}_i(h, oldsymbol{z_i}) \doteq rac{1}{oldsymbol{m}_i} \sum_{j=1}^{oldsymbol{m}_i} \ell(h(oldsymbol{x}_j), oldsymbol{y}_j)$$

Choose a malfare function such as a power mean:

$$M_p(i
ightarrow \mathcal{S}_i; oldsymbol{w}) \doteq \sqrt[p]{\sum_{i=1}^g} oldsymbol{w}_i \mathcal{S}_i^p$$

p = 1: **w**-weighted risk minimization $p \to \infty$: minimax fair learning

Our problem is *empirical malfare minimization*:

$$\hat{h} = \operatorname*{argmin}_{h \in \mathcal{H}} \Lambda \left(i \to \hat{R}_i(h); \boldsymbol{w} \right)$$



Rademacher Averages

- Rademacher averages $\mathfrak{R}_{m_i}(\ell \circ \mathcal{H}, \mathcal{D}_i)$ bound risk generalization gap
 - Suppose range r loss
 - Supremum Deviation (SD) Bound: With probability at least 1δ :

$$\forall i \in 1, \dots g : \sup_{h \in \mathcal{H}} \left| R_i(h) - \hat{R}_i(h) \right| \leq \varepsilon_i = 2 \hat{\mathbf{R}}_{\boldsymbol{m}_i}(\ell \circ \mathcal{H}, \boldsymbol{z}_i) + r \sqrt{\frac{\ln \frac{g}{\delta}}{2\boldsymbol{m}_i}}$$

• Can generalize this result to power-mean malfare

$$\sup_{h \in \mathcal{H}} \left| \mathcal{M}_p \left(i \mapsto \mathcal{R}_i(h); \boldsymbol{w} \right) \le \mathcal{M}_p \left(i \mapsto \hat{\mathcal{R}}_i(h); \boldsymbol{w} \right) \right| \le \max_{i \in 1, \dots, g} \boldsymbol{\varepsilon}_i$$

Theoretical Restricted Hypothesis Classes

Let $\varepsilon_i \doteq r\sqrt{\frac{\ln \frac{1}{\delta}}{2m_i}}$ and $\eta_i \doteq 2\mathfrak{R}_{m_i}(\ell \circ \mathcal{H}, \mathcal{D}_i) + \varepsilon_i$ We (pessimistically) upper-bound the objective value (w.h.p.)

$$\inf_{h' \in \mathcal{H}} \mathbf{M} \left(j \mapsto \hat{\mathbf{R}}(h', \boldsymbol{z}_j); \boldsymbol{w} \right) \leq \inf_{h' \in \mathcal{H}} \mathbf{M} \left(j \mapsto \begin{cases} j \neq i & \hat{\mathbf{R}}(h', \boldsymbol{z}_j) \\ j = i & \mathbf{R}(h', \mathcal{D}_i) + \boldsymbol{\varepsilon}_i \end{cases} \right)$$

and (optimistically) lower-bound the empirical malfare of all $h \in \mathcal{H}$ (w.h.p.)

$$M(j \mapsto \hat{R}(h, \boldsymbol{z}_j); \boldsymbol{w}) \ge M\left(j \mapsto \begin{cases} j \neq i & \hat{R}(h, \boldsymbol{z}_j) \\ j = i & R(h, \mathcal{D}_i) - \boldsymbol{\eta}_i \end{cases}\right)$$

Set $\mathcal{H}_i^* \doteq \{h \in \mathcal{H}\}$, where

$$M\left(j \mapsto \begin{cases} j \neq i & \hat{R}(h, z_j) \\ j = i & R(h, D_i) - \eta_i \end{cases}\right) \leq \inf_{h' \in \mathcal{H}} M\left(j \mapsto \begin{cases} j \neq i & \hat{R}(h', z_j) \\ j = i & R(h', D_i) + \varepsilon_i \end{cases}\right)$$

Theorem 1. Assume as above; the following then hold:

- 1. With probability at least $1-2\delta$ over choice of z_i , it holds that $\hat{h} \in \mathcal{H}_i^*$.
- 2. With probability at least $1-4\delta$ over choice of z_i ,

$$\left| \mathrm{R}(\hat{h}, \mathcal{D}_i) - \hat{\mathrm{R}}(\hat{h}, \boldsymbol{z}_i) \right| \leq 2 \mathfrak{R}_{\boldsymbol{m}_i} (\ell \circ \mathcal{H}_i^*, \mathcal{D}_i) + \boldsymbol{\varepsilon}_i$$
.

Empirical Restricted Hypothesis Classes

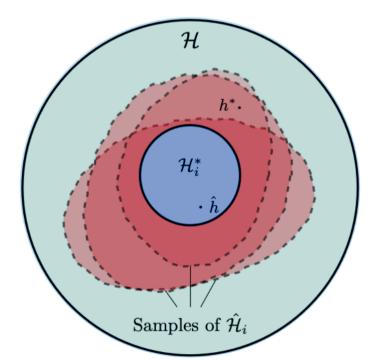
Take $\hat{\boldsymbol{\eta}}_i \doteq 2\hat{\mathbf{R}}_{\boldsymbol{m}_i}(\ell \circ \mathcal{H}, \boldsymbol{z}_i) + 2\boldsymbol{\varepsilon}_i$ Set $\hat{\mathcal{H}}_i \doteq \{h \in \mathcal{H}\}$, where

$$M\left(j \mapsto \begin{cases} j \neq i & \hat{R}(h, \boldsymbol{z}_j) \\ j = i & \hat{R}(h, \boldsymbol{z}_i) - \hat{\boldsymbol{\eta}}_i \end{cases}\right) \leq \inf_{h' \in \mathcal{H}} M\left(j \mapsto \begin{cases} j \neq i & \hat{R}(h', \boldsymbol{z}_j) \\ j = i & \hat{R}(h', \boldsymbol{z}_i) + 2\boldsymbol{\varepsilon}_i \end{cases}\right)$$

Theorem 1. Assume as above; the following then hold:

- 1. With probability at least $1 4\delta$ over choice of \mathbf{z}_i , it holds that $\hat{h} \in \mathcal{H}_i^* \subseteq \hat{\mathcal{H}}_i$.
- 2. With probability at least $1-6\delta$, it holds that

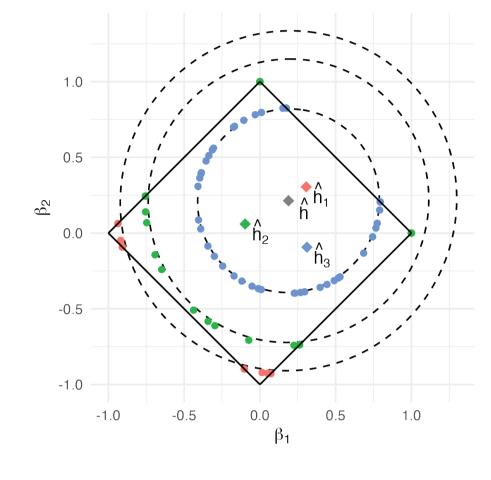
$$\left| \mathrm{R}(\hat{h}, \mathcal{D}_i) - \hat{\mathrm{R}}(\hat{h}, \boldsymbol{z}_i) \right| \leq 2 \hat{\mathbf{R}}_{\boldsymbol{m}_i} (\ell \circ \hat{\mathcal{H}}_i, \boldsymbol{z}_i) + 2 \boldsymbol{\varepsilon}_i .$$



Visualization of unrestricted class \mathcal{H} . theoretical restricted class \mathcal{H}_{i}^{*} , and samples of empirical restricted class $\hat{\mathcal{H}}_i$ (varying z_i).

Example with Linear Regression

- Take linear hypothesis class $oldsymbol{B} \ \doteq ig\{oldsymbol{eta} \in \mathbb{R}^2 : \|oldsymbol{eta}\|_1 \leq 1ig\}$
- Sample (\boldsymbol{x}_i, y_i) as $\boldsymbol{x}_i \sim \mathrm{Unif}([-1,1]^2)$ $y_i = \boldsymbol{x}_i \cdot \boldsymbol{\beta}_i + \mathrm{Unif}([-1,1])$
- Plot values of β which realize each supremum in the empirical Rademacher average for some Rademacher sample σ_k
- Points lie on either (the corner of) the ℓ_1 constraint boundary of \boldsymbol{B} or the restricted hypothesis constraint boundary of $\hat{\mathcal{H}}_i$



Group ◆ 1 ◆ 2 ◆ 3

Experiments with Logistic Regression

 $\mathcal{X} = [-1, 1]^{15}$

 $\mathcal{Y} = \pm 1$

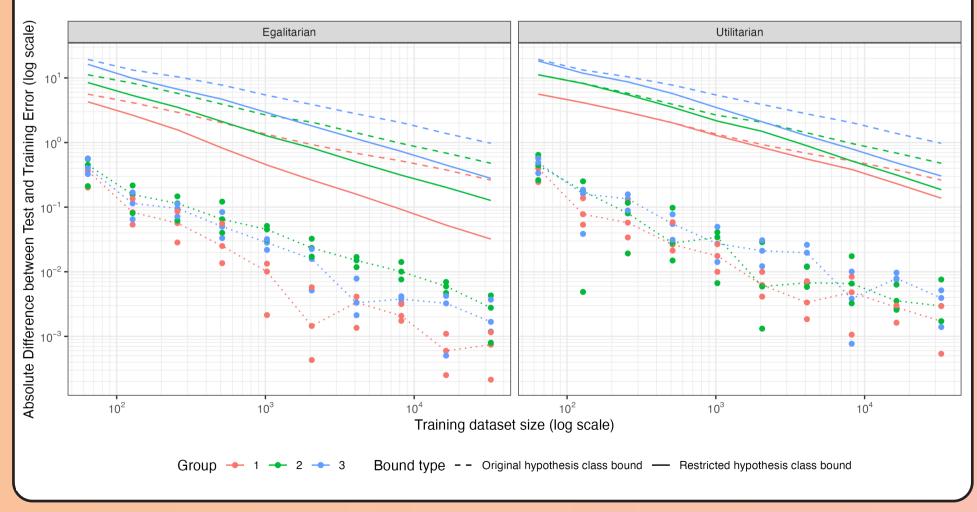
 $oldsymbol{x} \sim \mathrm{Unif}(\mathcal{X})$

 $\mathbb{P}(y=1) = \operatorname{logistic}(\boldsymbol{x} \cdot \boldsymbol{\beta}_i + \boldsymbol{\xi})$

$\hat{h} \doteq \operatorname*{argmin}_{h \in \mathcal{H}} \mathbf{M} (i \mapsto \hat{\mathbf{R}}(h, \boldsymbol{z}_i))$))
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$$egin{aligned} \mathcal{Y} &= \pm 1 \ egin{aligned} \mathbf{B} &\doteq \left\{oldsymbol{eta} \in \mathbb{R}^{15} \left| \left\| oldsymbol{eta}
ight\|_1 \leq 15
ight\} \end{aligned} \end{aligned} \qquad \hat{\mathrm{R}}(h, oldsymbol{z}_i) = rac{1}{oldsymbol{m}_i} \sum_{j=1}^{oldsymbol{m}_i} \mathrm{ln} ig(1 + \mathrm{exp}(oldsymbol{y}_{i,j} \cdot h(oldsymbol{x}_{i,j})) ig) \end{aligned}$$

	Data proportion	True parameters
Group 1	75%	$\beta_i = 0.3$
${\rm Group}\ 2$	$ \hspace{.08cm} 20\%$	$\beta_i = 0.1$
Group 3	5%	$\beta_i = 0.2$



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